Optimality and Sub-optimality of Principal Component Analysis for Spiked Random Matrices

> Alex Wein MIT

Joint work with: Amelia Perry (MIT), Afonso Bandeira (Courant NYU), Ankur Moitra (MIT)

Wigner Matrix

$$rac{1}{\sqrt{n}} W \in \mathbb{R}^{n imes n}$$
 symmetric,
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$$\frac{1}{\sqrt{n}}W \text{ Wigner, } \|x\| = 1.$$

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$$|eta| > \sqrt{\gamma}$$
, $\gamma = rac{n}{N}$, $eta \in [-1,\infty)$

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 (Wigner) vs $Y \sim \frac{1}{\sqrt{n}}W + \lambda x x^T$

• Detection: distinguish reliably (error prob \rightarrow 0)

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- Need a prior on the spike $x \in \mathbb{R}^n$
 - unit sphere
 - ▶ i.i.d. ±1
 - ▶ sparse ±1

Hypothesis testing power Recovery quality



Hypothesis testing power

Recovery quality



Hypothesis testing power Rec

Recovery quality



Hypothesis testing power Recov





This talk: focus on detection threshold (also hypothesis testing bounds, recovery threshold)

3 Scenarios

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2. Can beat PCA with an efficient algorithm (e.g. non-Gaussian Wigner)

3. Can beat PCA, but only with an inefficient algorithm (e.g. sparse priors; Wishart)

Sequence of distributions P_n is contiguous to Q_n if for any sequence of events A_n,

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► E.g. if
$$P_n$$
, Q_n have densities p_n , q_n :
 $\mathbb{E}_{Q_n} \left(\frac{dP_n}{dQ_n}\right)^2 = \mathbb{E}_{Y \sim Q_n} \left(\frac{p_n(Y)}{q_n(Y)}\right)^2$

L. Le Cam, 1960.

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$$\underbrace{\int_{A_n} dP_n}_{P_n(A_n)} = \int_{A_n} \frac{dP_n}{dQ_n} dQ_n \le \left(\underbrace{\int_{A_n} \left(\frac{dP_n}{dQ_n}\right)^2 dQ_n}_{\le \mathbb{E}_{Q_n}\left(\frac{dP_n}{dQ_n}\right)^2}\right)^{\frac{1}{2}} \left(\underbrace{\int_{A_n} dQ_n}_{Q_n(A_n)}\right)^{\frac{1}{2}}$$

► Taking
$$P_n : \frac{1}{\sqrt{n}}W + \lambda x x^T$$
 with $x \sim \mathcal{X}$, and $Q_n : \frac{1}{\sqrt{n}}W$

A. Montanari, D. Reichman, O. Zeitouni, NIPS 2015.

A. Onatski, M. J. Moreira, M. Hallin, AoS 2013.

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But what about when we know more about the spike?

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- Contiguity argument goes through for a general class of priors!
- But for sparse priors, with enough sparsity, PCA is no longer optimal.

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

$$\frac{1}{N}\sum_{k=1}^{N}y_{k}y_{k}^{T}$$

 $y_k \sim \mathcal{N}(0, I_n)$

VS

$$y_k \sim \mathcal{N}\left(0, I_n + \beta x x^T\right), x \sim \text{Unif}\left\{\pm \frac{1}{\sqrt{n}}\right\}^n$$

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Same story for spherical prior, contiguous for $|\beta| < \sqrt{\gamma}$ (PCA threshold).

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For Rademacher prior, something surprising happens:

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- If $\frac{n}{N} = \gamma < \frac{1}{3}$ then the models are contiguous for $|\beta| < \sqrt{\gamma}$
- But... for γ > 0.698 there exists a computationally inefficient procedure that distinguishes the two models for some β ∈ (−√γ, 0) (below the spectral threshold).

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Is there a computational gap?

Rademacher Wishart, Negative β



- PCA: succeeds above the line
- inefficient algorithm: succeeds above the line
- contiguity lower bound: impossible below the line

Back to Wigner: What if noise is not Gaussian?

$$Y = \frac{1}{\sqrt{n}} W + \lambda x x^T$$

 $x \sim \text{Unif}\{\mathbb{S}^{n-1}\}, \ W \in \mathbb{R}^{n \times n}$ but $W_{ij} \sim p(w)$ such that $\mathbb{E}w = 0, \ \mathbb{E}w^2 = 1.$

D. Feral, S. Peche, CMP 2006.

T. Tao, V. Vu, MAoRMT 2012.

Back to Wigner: What if noise is not Gaussian?

$$Y = \frac{1}{\sqrt{n}} W + \lambda x x^7$$

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Universality: spectral properties are unchanged...

D. Feral, S. Peche, CMP 2006.

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Can you tell which one is which? For $W_{ij} \sim \text{Unif}(\pm 1)$, $\lambda < 1$

 $W + \lambda \sqrt{n} x x^T$ vs W

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For $W_{ij} \sim \text{Unif}(\pm 1)$, $\lambda < 1$

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-1.0000	1.0000	-1.0000	-1.0000	1.0000	-1.0000
1.0000	1.0000	1.0000	-1.0000	-1.0000	1.0000
-1.0000	1.0000	1.0000	-1.0000	-1.0000	-1.0000
-1.0000	-1.0000	-1.0000	1.0000	-1.0000	1.0000
1.0000	-1.0000	-1.0000	-1.0000	-1.0000	1.0000
-1.0000	1.0000	-1.0000	1.0000	1.0000	1.0000

VS

-0.9988	1.0011	-1.0007	-0.9997	0.9990	-1.0014
1.0011	1.0010	0.9993	-0.9997	-1.0010	0.9987
-1.0007	0.9993	1.0004	-1.0002	-0.9994	-0.9991
-0.9997	-0.9997	-1.0002	1.0001	-1.0002	0.9997
0.9990	-1.0010	-0.9994	-1.0002	-0.9991	1.0012
-1.0014	0.9987	-0.9991	0.9997	1.0012	1.0017

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1.0000	1.0000	1.0000	-1.0000	-1.0000	1.0000
-1.0000	1.0000	1.0000	-1.0000	-1.0000	-1.0000
-1.0000	-1.0000	-1.0000	1.0000	-1.0000	1.0000
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Let's restrict ourselves to when the density p(w) is smooth.

If noise drawn from non-Gaussian p(w): we will beat PCA by applying some function $f : \mathbb{R} \to \mathbb{R}$ entrywise to our matrix $Y = W + \lambda \sqrt{nxx^{\top}}$, followed by PCA.

 $f(Y_{ij}) = f(W_{ij} + \lambda \sqrt{n} x_i x_j)$

$$\begin{array}{ll} f(Y_{ij}) &=& f\left(W_{ij} + \lambda \sqrt{n} x_i x_j\right) \\ &\approx& f\left(W_{ij}\right) + f'\left(W_{ij}\right) \lambda \sqrt{n} x_i x_j \end{array}$$

$$\begin{aligned} f(Y_{ij}) &= f\left(W_{ij} + \lambda\sqrt{n}x_ix_j\right) \\ &\approx f\left(W_{ij}\right) + f'\left(W_{ij}\right)\lambda\sqrt{n}x_ix_j \\ &\approx f\left(W_{ij}\right) + \mathbb{E}[f'\left(W_{ij}\right)]\lambda\sqrt{n}x_ix_j - \left(f'\left(W_{ij}\right) - \mathbb{E}f'\left(W_{ij}\right)\right)\lambda\sqrt{n}x_ix_j \end{aligned}$$

$$\begin{split} f(Y_{ij}) &= f\left(W_{ij} + \lambda \sqrt{n} x_i x_j\right) \\ &\approx f\left(W_{ij}\right) + f'\left(W_{ij}\right) \lambda \sqrt{n} x_i x_j \\ &\approx f\left(W_{ij}\right) + \mathbb{E}[f'\left(W_{ij}\right)] \lambda \sqrt{n} x_i x_j - \left(f'\left(W_{ij}\right) - \mathbb{E}f'\left(W_{ij}\right)\right) \lambda \sqrt{n} x_i x_j \\ &\approx f\left(W_{ij}\right) + \mathbb{E}[f'\left(W_{ij}\right)] \lambda \sqrt{n} x_i x_j. \end{split}$$

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$$\begin{aligned} f(Y_{ij}) &= f\left(W_{ij} + \lambda\sqrt{n}x_ix_j\right) \\ &\approx f\left(W_{ij}\right) + f'\left(W_{ij}\right)\lambda\sqrt{n}x_ix_j \\ &\approx f\left(W_{ij}\right) + \mathbb{E}[f'\left(W_{ij}\right)]\lambda\sqrt{n}x_ix_j - \left(f'\left(W_{ij}\right) - \mathbb{E}f'\left(W_{ij}\right)\right)\lambda\sqrt{n}x_ix_j \\ &\approx f\left(W_{ij}\right) + \mathbb{E}[f'\left(W_{ij}\right)]\lambda\sqrt{n}x_ix_j. \end{aligned}$$

It is (close to) a new spiked Wigner matrix with

$$\lambda' = \frac{\lambda \mathbb{E} f'(W_{ij})}{\sqrt{\mathbb{E} f^2(W_{ij})}}.$$

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Calculus of variations gives optimal choice of f:

$$f(w) = \frac{-p'(w)}{p(w)}$$



Figure: Dashed: p(w), Solid: f(w) = -p'(w)/p(w)

T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015.

F. Krzakala, J. Xu, L. Zdeborová, 2016.





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F. Krzakala, J. Xu, L. Zdeborová, 2016.

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• New threshold at
$$\lambda = \frac{1}{\sqrt{F_{\rho}}}, \quad F_{\rho} = \mathbb{E}_{\rho} \left(\frac{p'(w)}{\rho(w)} \right)^2 \ge 1.$$

F. Krzakala, J. Xu, L. Zdeborová, 2016.

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Contiguity shows this is optimal!

- T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015.
- F. Krzakala, J. Xu, L. Zdeborová, 2016.

$$\mathbb{E}_{Q_n}\left(\frac{dP_n}{dQ_n}\right)^2 = \mathbb{E}\exp\left(\frac{\lambda^2 n}{2}\langle x, x'\rangle^2\right) \qquad x, x' \sim \mathcal{X}$$

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How big a parabola can you fit underneath the rate function?

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- How big a parabola can you fit underneath the rate function?
- ► E.g. Rademacher prior (±1) has $R(u) = \log 2 h\left(\frac{1+u}{2}\right)$ where $h(p) = -p \log p - (1-p) \log(1-p)$



J. Banks, C. Moore, J. Neeman, P. Netrapalli, COLT 2016.

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Condition away from rare bad events

J. Banks, C. Moore, J. Neeman, P. Netrapalli, COLT 2016.

• Sparsity $\rho \in [0, 1]$

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• P_n : spiked Wigner with prior \mathcal{X}_{ρ} , Q_n : unspiked

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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- *P̃_n*: change prior to *X̃_ρ*: condition on close-to-typical proportion of nonzeros

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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Sparse Rademacher: Results



- unconditioned
- conditioned
- noise-conditioned (upcoming)
- replica prediction (truth)

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Thanks! Questions?